

Lecture #2: Friday, 7 Feb, 2025  
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## 1 Khintchine's inequality

We aim to prove the matrix Khintchine's inequality. But before that, let's look at the same for scalar random variables.

### Theorem 1

Let  $a_1, \dots, a_n \in \mathbb{R}$  and  $t \in \mathbb{N}$ . Then

$$\mathbb{E} \left[ \left( \sum_{i=1}^n g_i a_i \right)^{2t} \right]^{\frac{1}{2t}} \leq \sqrt{2t-1} \sqrt{\sum_{i=1}^n a_i^2}$$

where the expectation is taken over independent Gaussians  $g_1, \dots, g_n \sim \mathcal{N}(0, 1)$ .

*Proof.* Denote  $\mathbf{a} = (a_1, \dots, a_n)$ ,  $\mathbf{g} = (g_1, \dots, g_n)$ . Observe that  $g := \frac{\langle \mathbf{g}, \mathbf{a} \rangle}{\|\mathbf{a}\|_2} \sim \mathcal{N}(0, 1)$  if  $g_i$ 's are independent standard normals. Recall that  $\mathbb{E}[g^{2t}] = \prod_{i=1}^t (2i-1) \leq (2t-1)^t$  for such  $g$ . The former equality has been proven in Appendix A.1. Multiplying throughout by  $\|\mathbf{a}\|_2^{2t}$  and taking  $(2t)^{\text{th}}$  square roots gives the desired result. ■

We now move on to Khintchine's inequality for matrices whose proof relies on a bound stated in Lemma 10 in Appendix A.2.

### Theorem 2

$\exists c > 0$  such that  $\forall$  symmetric  $A_1, \dots, A_n \in \mathbb{R}^{m \times m}$  (with  $m \geq 3$ ) we have

$$\mathbb{E} \left[ \left\| \sum_{i=1}^n g_i A_i \right\|_2 \right] \leq c \sqrt{\lg m} \left\| \sum_{i=1}^n A_i^2 \right\|_2^{\frac{1}{2}}$$

where the expectation is taken over independent Gaussians  $g_1, \dots, g_n \sim \mathcal{N}(0, 1)$ .

*Proof.* Denote  $X := \sum_{i=1}^n g_i A_i$ . Then  $X$  is symmetric and  $\|X\|_2$  is the absolute value of its largest sized eigenvalue. Say  $X$  has eigenvalues  $\lambda_1, \dots, \lambda_m \in \mathbb{R}$  such that  $|\lambda_m| \geq |\lambda_i| \forall i \in [m]$ . Let  $\sigma = \left\| \sum_{i=1}^n A_i^2 \right\|_2$  which is also the absolute value of the eigenvalue of  $\sum A_i^2$  with largest size. Then for any  $t \in \mathbb{N}$  we have  $\|X\|_2^{2t} = |\lambda_m|^{2t} \leq \sum_{i=1}^m \lambda_i^{2t} = \text{Tr}(X^{2t})$ , whence

$$\begin{aligned} \mathbb{E} [\|X\|_2] &\leq \mathbb{E} \left[ (\text{Tr}(X^{2t}))^{\frac{1}{2t}} \right] \\ &\stackrel{\text{Jensen}}{\leq} \left( \mathbb{E} [\text{Tr}(X^{2t})] \right)^{\frac{1}{2t}} \\ &\stackrel{\text{Lemma 10}}{\leq} \sqrt{2t-1} \left[ \text{Tr} \left[ \left( \sum_{i=1}^n A_i^2 \right)^t \right] \right]^{\frac{1}{2t}} \\ &\leq \sqrt{2t-1} (m\sigma^{2t})^{\frac{1}{2t}} = \sqrt{2t-1} \sigma m^{\frac{1}{2t}}. \end{aligned}$$

Choosing  $t \in \left[ \frac{\lg m - 1}{2}, \frac{\lg m + 1}{2} \right] \cap \mathbb{Z}$  gives  $\mathbb{E} [\|X\|_2] \leq \sigma \sqrt{\lg m} \cdot \underbrace{2^{\frac{1}{2} + \frac{1}{\lg(9/4)}}}_c$ . ■

Now let's move on to Khintchine's inequality with Rademacher weights, instead of Gaussian. We expect such a statement to be true because somehow the 'tails' in the Rademacher case are bounded above by Gaussians. What we precisely want is stated in Appendix A.3.

### Theorem 3

$\exists c > 0$  such that  $\forall$  symmetric  $A_1, \dots, A_n \in \mathbb{R}^{m \times m}$ , we have

$$\mathbb{E} \left[ \left\| \sum_{i=1}^n b_i A_i \right\|_2 \right] \leq c \sqrt{\lg m} \left\| \sum_{i=1}^n A_i^2 \right\|_2^{\frac{1}{2}}$$

where the expectation is taken over a uniform draw of  $\mathbf{b} \in \{\pm 1\}^n$ .

*Proof.* Immediate from Theorem 2 and Lemma 11 (different  $c$  from Theorem 2). ■

## 2 Sums of independent random variables

We now want to look at quantities like  $\left\| \sum_{i=1}^n (H_i - \mathbb{E}[H_i]) \right\|_2$  where the  $H_i$ 's are chosen uniformly from the set of  $m \times m$  real symmetric matrices. In such cases, a symmetrization

trick is really useful which is presented in Appendix A.4. We shall use this to prove the following expected deviation.

**Theorem 4**

$\exists c > 0$  such that if  $H_1, \dots, H_n \in \mathbb{R}^{m \times m}$  are symmetric and chosen uniformly randomly, we have

$$\mathbb{E}_{\mathbf{H}} \left[ \left\| \sum_{i=1}^n (H_i - \mathbb{E}[H_i]) \right\|_2 \right] \leq c \sqrt{\lg m} \mathbb{E}_{\mathbf{H}} \left[ \left\| \sum_i H_i^2 \right\|_2^{\frac{1}{2}} \right].$$

*Proof.* Lemma 12 says that the LHS is at most  $2 \mathbb{E}_{\mathbf{H}} \left[ \left\| \sum_{i=1}^n b_i H_i \right\|_2 \right]$ . But this is the same

as  $2 \mathbb{E}_{\mathbf{H}} \left[ \mathbb{E}_{\mathbf{b} \in \{\pm 1\}^n} \left( \left\| \sum_{i=1}^n b_i H_i \right\|_2 \mid H_1, \dots, H_n \right) \right]$  by the law of total expectation. By Theo-

rem 3, this expression is at most  $\underbrace{2c'}_c \sqrt{\lg m} \mathbb{E}_{\mathbf{H}} \left[ \mathbb{E}_{\mathbf{b} \in \{\pm 1\}^n} \left( \left\| \sum_{i=1}^n H_i^2 \right\|_2^{\frac{1}{2}} \mid H_1, \dots, H_n \right) \right] =$

$$c \sqrt{\lg m} \mathbb{E}_{\mathbf{H}} \left[ \left\| \sum_{i=1}^n H_i^2 \right\|_2^{\frac{1}{2}} \right]. \quad \blacksquare$$

## 2.1 Chernoff bound for random symmetric matrices

**Theorem 5**

$\exists c, c' > 0$  such that if  $H_1, \dots, H_n \in \mathbb{R}^{m \times m}$  are symmetric and chosen uniform and independently such that  $H_i \succeq 0$  a.s. and  $H_i \preceq M\mathbb{I}$  a.s., we have the following  $\forall \varepsilon > 0$ :

$$\mathbb{E} \left[ \left\| \sum_{i=1}^n H_i \right\|_2 \right] \leq (1 + \varepsilon) \left\| \sum \mathbb{E}[H_i] \right\|_2 + \left( 1 + \frac{1}{\varepsilon} \right) cM \sqrt{\lg m}$$

$$\stackrel{\text{if } \lambda_{\max}(\sum \mathbb{E}[H_i]) \gg M \lg m}{\leq} (1 + \mathcal{O}(\varepsilon)) \left\| \sum_{i=1}^n \mathbb{E}[H_i] \right\|_2.$$

*Proof.* First write  $\mathbb{E} \left[ \left\| \sum_{i=1}^n H_i \right\|_2 \right] \leq \mathbb{E} \left[ \left\| \sum_i H_i - \mathbb{E}[H_i] \right\|_2 \right] + \left\| \sum_i \mathbb{E}[H_i] \right\|_2$  by the tri-

angle inequality. Then  $\mathbb{E} \left[ \left\| \sum_i H_i - \mathbb{E}[H_i] \right\|_2 \right] \stackrel{\text{Theorem 4}}{\lesssim} \sqrt{\lg m} \mathbb{E} \left[ \left\| \sum_i H_i^2 \right\|_2^{\frac{1}{2}} \right] \stackrel{0 \preceq H_i \preceq M\mathbb{I}}{\lesssim}$

$$\sqrt{\lg m} \cdot \mathbb{E} \left[ \sqrt{M} \left\| \sum_i H_i \right\|_2^{\frac{1}{2}} \right] \leq \sqrt{M \lg m} \mathbb{E} \left[ \left\| \sum_i H_i \right\|_2 \right]^{\frac{1}{2}}. \text{ Taking } x := \sqrt{\mathbb{E} \left[ \left\| \sum_{i=1}^n H_i \right\|_2 \right]}$$

gives from the last inequality that  $x^2 - xc\sqrt{M \log m} - \|\sum \mathbb{E}[H_i]\|_2 \leq 0$ . Since this is a (strictly) convex quadratic, we must have  $x \leq \frac{c\sqrt{M \log m} + \sqrt{c^2 M \log m + 4 \|\sum \mathbb{E}[H_i]\|_2}}{2}$  for the previous inequality to be true. This in turn implies that

$$\begin{aligned} \mathbb{E} \left[ \left\| \sum H_i \right\|_2 \right] &= x^2 \leq \frac{1}{2} c^2 M \log m + \left\| \sum \mathbb{E}[H_i] \right\|_2 + \\ &\quad \frac{1}{2} \sqrt{\frac{1}{\varepsilon} c^2 M \log m} \cdot \sqrt{\varepsilon} \sqrt{c^2 M \log m + 4 \left\| \sum \mathbb{E}[H_i] \right\|_2} \\ &\stackrel{\text{GM} \leq \text{AM}}{\leq} \frac{1}{2} c^2 M \log m + \left\| \sum \mathbb{E}[H_i] \right\|_2 + \\ &\quad \frac{\frac{1}{\varepsilon} c^2 M \log m + \varepsilon c^2 M \log m + 4\varepsilon \left\| \sum \mathbb{E}[H_i] \right\|_2}{4} \\ &= \frac{1}{4} \left( 2 + \varepsilon + \frac{1}{\varepsilon} \right) c^2 M \log m + (1 + \varepsilon) \left\| \sum \mathbb{E}[H_i] \right\|_2. \end{aligned}$$

The above is true for any  $\varepsilon > 0$ . Doing the same argument with  $\varepsilon$  replaced by  $\frac{1}{\varepsilon}$  gives a coefficient of  $(1 + \frac{1}{\varepsilon})$  in the second term (first term is unchanged), hence we can assume  $\varepsilon < 1$ . Thus, the above inequality for  $\varepsilon \in (0, 1)$  can be stated as

$$\mathbb{E} \left[ \left\| \sum H_i \right\|_2 \right] \leq \left( 1 + \frac{1}{\varepsilon} \right) c' M \log m + (1 + \varepsilon) \left\| \sum \mathbb{E}[H_i] \right\|_2.$$

■

## 2.2 Bernstein bound for random symmetric matrices

### Theorem 6

$\exists c_1, c_2 > 0$  such that if  $H_1, \dots, H_n \in \mathbb{R}^{m \times m}$  are chosen uniformly randomly from the space of real symmetric matrices with  $\mathbb{E}[H_i] = 0$  and  $\|H_i\|_2 \leq M$  almost surely  $\forall i$ , then

$$\mathbb{E} \left[ \left\| \sum H_i \right\|_2 \right] \leq c_1 \sqrt{\log m} \left\| \sum \mathbb{E}[H_i^2] \right\|_2^{\frac{1}{2}} + c_2 M \log m.$$

*Proof.*  $\mathbb{E} \left[ \left\| \sum H_i \right\|_2 \right] \stackrel{\text{Theorem 4}}{\leq} c \sqrt{\log m} \mathbb{E} \left[ \left\| \sum H_i^2 \right\|_2^{\frac{1}{2}} \right] \stackrel{\text{Jensen}}{\leq} c \sqrt{\log m} \mathbb{E} \left[ \left\| \sum H_i^2 \right\|_2 \right]^{\frac{1}{2}} \stackrel{\text{Theorem 5}}{\leq} c \sqrt{\log m} \left( c' \left\| \sum \mathbb{E}[H_i^2] \right\|_2^{\frac{1}{2}} + c'' M \sqrt{\log m} \right)$  where the last inequality is true by replacing  $H_i, M$  with  $H_i^2, M^2$  in Theorem 5 and then applying the inequality  $\sqrt{a+b} \leq \sqrt{a} + \sqrt{b}$  for  $a, b \geq 0$  and end by taking  $c_1 := cc', c_2 := cc''$ . ■

### 3 Applications

#### 3.1 Another proof of Graph Sparsification

**Problem statement and algorithm.** Let's recall the problem of graph sparsification. Given an unweighted undirected graph  $G = (V, E)$  with  $n = |V|$  we want to find another undirected *weighted* graph  $G'(V', E', w)$  with  $E' \subseteq E$  such that (with multiplicative error)  $\sum_{e \in E'} w_e L_e \stackrel{1 \pm \varepsilon}{\simeq} \sum_{e \in E} L_e$  (so that cut-values are preserved), where the edge Laplacians are  $L_{\{i,j\}} := \delta_{ii} + \delta_{jj} - \delta_{ij} - \delta_{ji}$  where  $\delta_{\alpha\beta}$  is the  $n \times n$  matrix with a 1 entry in position  $(\alpha, \beta)$  and zero elsewhere. Recall the Laplacian of the graph  $(L =) L_G = \sum_{e \in E} L_e$  and the normalized edge Laplacian  $\tilde{L}_e = L^{\dagger/2} L_e L^{\dagger/2}$ . With this new notation, our goal is the same as  $\sum_{e \in E'} w_e \tilde{L}_e \stackrel{1 \pm \varepsilon}{\simeq} \mathbf{I}_{\perp}$ . Look at [Lecture 1](#) for details.

The ideal algorithm was based on importance sampling, to independently sample edge  $e \in E$  with probability  $p_e := \frac{1}{R} \|\tilde{L}_e\|_2$  to be included in  $E'$ , where  $R$  is to be chosen later.

If  $e \in E'$  we will set  $w_e = \frac{1}{p_e}$ . Then  $\mathbb{E}[|E'|] = \sum_{e \in E} p_e = \sum_{e \in E} \frac{1}{R} \|\tilde{L}_e\|_2 \stackrel{(*)}{=} \sum_{e \in E} \frac{1}{R} \text{Tr}(\tilde{L}_e) = \frac{1}{R} \text{Tr}(\mathbf{I}_{\perp}) = \frac{n-1}{R}$ . (\*) follows from the fact that each  $\tilde{L}_e \succeq 0$  and has rank 1. To prove that  $|E'|$  is concentrated around this number, we can use the usual Chernoff bound.

**Error analysis.** Let  $Q_e := \begin{cases} \frac{1}{p_e} \tilde{L}_e & \text{if } e \in E' \\ \mathbf{0}_{n \times n} & \text{otherwise} \end{cases} = \underbrace{\mathbf{1}[e \text{ sampled}]}_{w_e} \cdot \frac{1}{p_e} \tilde{L}_e$  for each  $e \in E$  (if

an edge is not present in  $E'$ , then think of its weight as 0). Each  $Q_e$  is positive semidefinite and has all eigenvalues  $\leq R$ . So  $\mathbb{E}[Q_e] = \tilde{L}_e$ . So  $\mathbb{E} \left[ \left\| \sum_{e \in E'} w_e \tilde{L}_e \right\|_2 \right] = \mathbb{E} \left[ \left\| \sum_{e \in E} Q_e \right\|_2 \right] \leq$

$(1+\varepsilon) \left\| \sum_{e \in E} \mathbb{E}[Q_e] \right\|_2 + \left(1 + \frac{1}{\varepsilon}\right) cR \log n$  where the last inequality follows from Theorem 5.

Choosing  $R = \frac{\varepsilon^2}{\log n}$  and noting that  $\left\| \sum_{e \in E} \mathbb{E}[Q_e] \right\|_2 = \left\| \sum_{e \in E} \tilde{L}_e \right\|_2 = \|\mathbf{I}_{\perp}\|_2 = 1$  guarantees

$$\mathbb{E} \left[ \left\| \sum_{e \in E'} w_e \tilde{L}_e \right\|_2 \right] \leq 1 + \mathcal{O}(\varepsilon).$$

#### 3.2 ‘Machine Learning’

We have an (unknown) distribution  $D$  on  $R^d$  with a random variable  $X \sim D$  satisfying  $\mathbb{E}_D[X] = 0, \mathbb{E}_D[XX^T] = \Sigma$  ( $\Sigma$  also unknown). To *learn*  $\Sigma$ , an intuitive algorithm is

to sample  $X_1, \dots, X_m \sim D$  independently and report an estimate  $\hat{\Sigma} := \frac{1}{m} \sum_{i=1}^m X_i X_i^\top$ . We want to get theoretical guarantees of the form “ $\|\hat{\Sigma} - \Sigma\|_2 \leq \varepsilon$  with probability 99%”. For ‘real life’ cases, we assume  $\Sigma$  is invertible. Since the above estimate can depend on the eigenvalues of  $\Sigma$ , we instead demand a relative estimate so that people can have an overall idea of  $\varepsilon$ ’s are good enough. So we instead demand a guarantee of the form “ $\|\Sigma^{-1/2} \hat{\Sigma} \Sigma^{-1/2} - \mathbf{I}\|_2 \leq \varepsilon$  with probability 99%”. This is equivalent to demanding “ $\mathbf{v}^\top (\Sigma^{-1/2} \hat{\Sigma} \Sigma^{-1/2}) \mathbf{v} \in (1 - \varepsilon, 1 + \varepsilon) \|\mathbf{v}\|^2 \forall \mathbf{v} \in \mathbb{R}^d$  with probability 99%”.

Comparing Theorem 6 with the above expression forces us to write it as  $\Sigma^{-1/2} \hat{\Sigma} \Sigma^{-1/2} = \frac{1}{m} \sum_{i=1}^m Y_i Y_i^\top$  where  $Y_i := \Sigma^{-1/2} X_i$ . Noting that  $\mathbb{E}[Y_i Y_i^\top] = \Sigma^{-1/2} \Sigma \Sigma^{-1/2} = \mathbf{I}$  suggests that we can use Theorem 6 with  $H_i := \frac{1}{m} (Y_i Y_i^\top - \mathbf{I})$ , provided we know some  $K$  such that  $K \geq \|\Sigma^{-1/2} X\|_2$  almost surely. The latter assumption guarantees  $\|H_i\| \leq \frac{1}{m} (\|Y_i\|^2 + 1) \lesssim \frac{K^2}{m}$ . Using the raw form of Bernouli (Theorem 4) gives  $\mathbb{E} \left[ \left\| \sum H_i \right\|_2 \right] \leq c \sqrt{\log d} \cdot \frac{K^2}{m} \sqrt{m} = cK^2 \sqrt{\frac{\log d}{m}}$ . Choosing  $m > \frac{c^2 K^4 \log d}{\varepsilon^2}$  samples does the job.

## A Appendix

### A.1 Even moments of Gaussians

**Proposition 7** (Gaussian integration by parts)

If  $X \sim \mathcal{N}(0, 1)$  and  $f$  is an differentiable function such that its derivative is integrable then  $\mathbb{E}[Xf(X)] = \mathbb{E}[f'(X)]$ .

**Lemma 8** (even Gaussian moments)

If  $X \sim \mathcal{N}(0, 1)$  and  $t \in \mathbb{N}$  then  $\mathbb{E}[X^{2t}] = \prod_{i=1}^t (2i - 1)$ .

*Proof.* Taking  $f(x) = x^{2t-1}$  in Proposition 7 gives  $\mathbb{E}[X^{2t}] = \mathbb{E}[X \cdot f(X)] = (2t - 1)\mathbb{E}[X^{2t-2}]$ . This is one step in our calculation. Doing this recursively gives  $\mathbb{E}[X^{2t}] = \prod_{i=2}^t (2i - 1) \cdot \mathbb{E}[X^2] = \prod_{i=1}^t (2i - 1)$  because  $\mathbb{E}[X^2] = \text{Var}[X] = 1$ . ■

### A.2 Bounding expectation of trace of even powers of Gaussian linear combination of matrices

**Lemma 9**

Let  $X, A_1, \dots, A_n \in \mathbb{R}^{m \times m}$  be symmetric matrices,  $s \in \mathbb{N}$ . Then  $\text{Tr}(A_k X^l A_k X^{2s-l}) \leq \text{Tr}(A_k^2 X^{2s})$  for every  $k \in [n], l \in [2s] \cup \{0\}$ .

*Proof.* Consider the function  $f(k, l) = \text{Tr}(A_k X^l A_k X^{2s-l})$  for each  $k \in [n], l \in [2s] \cup \{0\}$ .  $X$  is symmetric, hence it can be decomposed as  $X = \sum_{i=1}^m \lambda_i \mathbf{v}_i \mathbf{v}_i^\top$  where  $\{\mathbf{v}_i\}_{i=1}^m$  is an orthonormal basis of  $\mathbb{R}^m$ , and  $\lambda_i \in \mathbb{R}$  are the eigenvalues of  $X$ . Then

$$\begin{aligned} f(k, l) &= \text{Tr}(A_k X^l A_k X^{2s-l}) \\ &= \text{Tr} \left( A_k \left( \sum_{i=1}^m \lambda_i \mathbf{v}_i \mathbf{v}_i^\top \right)^l A_k \left( \sum_{j=1}^m \lambda_j \mathbf{v}_j \mathbf{v}_j^\top \right)^{2s-l} \right) \\ &= \text{Tr} \left( A_k \left( \sum_{i=1}^m \lambda_i^l \mathbf{v}_i \mathbf{v}_i^\top \right) A_k \left( \sum_{j=1}^m \lambda_j^{2s-l} \mathbf{v}_j \mathbf{v}_j^\top \right) \right) \\ &= \sum_{i,j} \lambda_i^l \lambda_j^{2s-l} \text{Tr} (A_k \mathbf{v}_i \mathbf{v}_i^\top A_k \mathbf{v}_j \mathbf{v}_j^\top) \\ &\leq \sum_{i,j} |\lambda_i|^l |\lambda_j|^{2s-l} \text{Tr} (A_k \mathbf{v}_i \mathbf{v}_i^\top A_k \mathbf{v}_j \mathbf{v}_j^\top) \end{aligned}$$

The expression  $\sum_{i,j} |\lambda_i|^l |\lambda_j|^{2s-l} \text{Tr}(A_k \mathbf{v}_i \mathbf{v}_i^\top A_k \mathbf{v}_j \mathbf{v}_j^\top)$  is convex in  $l$ . Hence it maximizes at the endpoints of the range of  $l$ , namely  $0, 2s$ . Due to the symmetry of  $A_k$ , the value of the expression at  $l = 0$  is the same as that at  $l = 2s$ . This common value is  $\sum_{i,j} \lambda_j^{2s} \text{Tr}(A_k \mathbf{v}_i \mathbf{v}_i^\top A_k \mathbf{v}_j \mathbf{v}_j^\top)$  which is the same as  $f(k, 0)$  (this step would not be possible if  $2s$  were replaced by some odd number). Thus we have proven that for any  $k \in [n], l \in [2s] \cup \{0\}$ ,  $\text{Tr}(A_k X^l A_k X^{2s-l}) = f(k, l) \leq f(k, 0) = \text{Tr}(A_k^2 X^{2s})$ . ■

### Lemma 10

Let  $A_1, \dots, A_n \in \mathbb{R}^{n \times n}$  be symmetric matrices and  $g_1, \dots, g_n \sim \mathcal{N}(0, 1)$  be independent

Gaussians. Denote  $X := \sum_{i=1}^n g_i A_i$ . Then

$$\left(\mathbb{E}[\text{Tr}(X^{2t})]\right)^{\frac{1}{2t}} \leq \sqrt{2t-1} \left[ \text{Tr} \left[ \left( \sum_{i=1}^n A_i^2 \right)^t \right] \right]^{\frac{1}{2t}}.$$

*Proof.* We will first use Proposition 7 similar to the proof of Lemma 8, and then invoke Lemma 9 with  $s = t - 1$ . We have

$$\begin{aligned} \mathbb{E}[\text{Tr}(X^{2t})] &= \mathbb{E}[\text{Tr}(X \cdot X^{2t-1})] \\ &= \sum_{k=1}^n \mathbb{E}[g_k \text{Tr}(A_k X^{2t-1})] \\ &\stackrel{\text{Proposition 7}}{=} \sum_{k=1}^n \sum_{l=0}^{2t-2} \mathbb{E} \left[ \text{Tr} \left( A_k X^l \underbrace{(\partial_{g_k} X)}_{A_k} X^{2t-2-l} \right) \right] \\ &= \sum_{k=1}^n \sum_{l=0}^{2t-2} \mathbb{E}[\text{Tr}(A_k X^l A_k X^{2t-2-l})] \\ &\stackrel{\text{Lemma 9}}{\leq} \sum_{l=0}^{2t-2} \sum_{k=1}^n \mathbb{E}[\text{Tr}(A_k^2 X^{2t-2})] \\ &= \sum_{l=0}^{2t-2} \mathbb{E} \left[ \text{Tr} \left( \left( \sum_{k=1}^n A_k^2 \right) X^{2t-2} \right) \right] \\ &= (2t-1) \mathbb{E} \left[ \text{Tr} \left( \left( \sum_{k=1}^n A_k^2 \right) X^{2t-2} \right) \right] \\ &\stackrel{\text{Hölder}}{\leq} (2t-1) \mathbb{E} \left[ \left[ \text{Tr} \left( \left( \sum_{k=1}^n A_k^2 \right)^t \right) \right]^{\frac{1}{t}} \cdot (\text{Tr}(X^{2t}))^{1-\frac{1}{t}} \right] \end{aligned}$$



$$\begin{aligned}
 &= (2t - 1) \left[ \text{Tr} \left( \left( \sum_k A_k^2 \right)^t \right) \right]^{\frac{1}{t}} \mathbb{E} \left[ (\text{Tr}(X^{2t}))^{1-\frac{1}{t}} \right] \\
 &\stackrel{\text{Jensen}}{\leq} (2t - 1) \left[ \text{Tr} \left( \left( \sum_k A_k^2 \right)^t \right) \right]^{\frac{1}{t}} (\mathbb{E} [\text{Tr}(X^{2t})])^{1-\frac{1}{t}} \\
 \implies &(\mathbb{E} [\text{Tr}(X^{2t})])^{\frac{1}{t}} \leq (2t - 1) \left[ \text{Tr} \left( \left( \sum_k A_k^2 \right)^t \right) \right]^{\frac{1}{t}} \\
 \implies &(\mathbb{E} [\text{Tr}(X^{2t})])^{\frac{1}{2t}} \leq \sqrt{2t - 1} \left[ \text{Tr} \left( \left( \sum_k A_k^2 \right)^t \right) \right]^{\frac{1}{2t}}.
 \end{aligned}$$

■

### A.3 Gauss dominates Rademacher

#### Lemma 11

If  $\mathbf{b}$  is uniformly random on  $\{\pm 1\}^n$  and  $g_1, \dots, g_n \sim \mathcal{N}(0, 1)$  are independent (and also independent of  $\mathbf{b}$ ), then

$$\mathbb{E}_{\mathbf{b}} \left\| \sum_i b_i A_i \right\|_2 \leq \sqrt{\frac{\pi}{2}} \mathbb{E}_{\mathbf{g}} \left\| \sum_i g_i A_i \right\|_2.$$

*Proof.*  $\mathbb{E}_{\mathbf{g}} \left\| \sum_i g_i A_i \right\|_2 = \mathbb{E}_{\mathbf{g}, \mathbf{b}} \left\| \sum_i |g_i| b_i A_i \right\|_2 \geq \mathbb{E}_{\mathbf{b}} \left\| \sum_i b_i \mathbb{E}_{\mathbf{g}} |g_i| A_i \right\|_2 = \sqrt{\frac{2}{\pi}} \mathbb{E}_{\mathbf{b}} \left\| \sum_i b_i A_i \right\|_2. \quad \blacksquare$

### A.4 Symmetrization trick with Rademacher

#### Lemma 12

If  $H_1, \dots, H_n \in \mathbb{R}^{m \times m}$  are chosen uniformly randomly and independently from the set of  $m \times m$  real symmetric matrices, then

$$\mathbb{E}_{\mathbf{H}} \left[ \left\| \sum_{i=1}^n (H_i - \mathbb{E}[H_i]) \right\|_2 \right] \leq 2 \mathbb{E}_{\mathbf{b} \in \{\pm 1\}^n} \left[ \left\| \sum_{i=1}^n b_i H_i \right\|_2 \right].$$

*Proof.* The symmetrization trick is to introduce an independent identical copy of each  $H_i$ , namely  $H'_i$ , then each  $H_i - \mathbb{E}[H_i]$  becomes  $\mathbb{E}_{\mathbf{H}'} [H_i - H'_i]$  because  $\mathbb{E}[H_i] = \mathbb{E}[H'_i]$ . Then starting from the LHS we have,

$$\mathbb{E}_{\mathbf{H}} \left[ \left\| \sum_{i=1}^n (H_i - \mathbb{E}[H_i]) \right\|_2 \right] = \mathbb{E}_{\mathbf{H}} \left[ \left\| \mathbb{E}_{\mathbf{H}'} \left[ \sum_{i=1}^n H_i - H'_i \right] \right\|_2 \right]$$

$$\begin{aligned}
& \stackrel{\text{Jensen}}{\leq} \mathbb{E} \mathbb{E}_{\mathbf{H}, \mathbf{H}'} \left[ \left\| \sum_i H_i - H'_i \right\|_2 \right] \\
& \stackrel{(*)}{=} \mathbb{E}_{\mathbf{H}, \mathbf{H}'} \left[ \mathbb{E}_{\mathbf{b} \in \{\pm 1\}^n} \left\| \sum_{i=1}^n b_i (H_i - H'_i) \right\|_2 \right] \\
& \stackrel{(**)}{\leq} \mathbb{E}_{\mathbf{H}, \mathbf{H}'} \left[ \mathbb{E}_{\mathbf{b} \in \{\pm 1\}^n} \left\| \sum_{i=1}^n b_i H_i \right\|_2 \right] + \mathbb{E}_{\mathbf{H}, \mathbf{H}'} \left[ \mathbb{E}_{\mathbf{b} \in \{\pm 1\}^n} \left\| \sum_{i=1}^n b_i H'_i \right\|_2 \right] \\
& = 2 \mathbb{E}_{\mathbf{H}} \left\| \sum_{i=1}^n b_i H_i \right\|_2.
\end{aligned}$$

$(*)$  holds because  $H_i - H'_i$  has a symmetric distribution and hence  $b_i(H_i - H'_i) \stackrel{D}{=} H_i - H'_i$ .

$(**)$   $\leq$  is true by triangle inequality on  $\|\cdot\|_2$ . ■