

## Lecture 14: SVD, Power method, and Planted Graph problems (+ eigenvalues of random matrices)

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Scribe:

Today we continue the topic of low-dimensional approximation to datasets and matrices. Last time we saw the singular value decomposition of matrices.

## 1 SVD computation

Recall this theorem from last time.

**THEOREM 1 (SINGULAR VALUE DECOMPOSITION AND BEST RANK- $k$ -APPROXIMATION)**

An  $m \times n$  real matrix has  $t \leq \min\{m, n\}$  nonnegative real numbers  $\sigma_1, \sigma_2, \dots, \sigma_t$  (called singular values) and two sets of unit vectors  $U = \{u_1, u_2, \dots, u_t\}$  which are in  $\mathbb{R}^m$  and  $V = \{v_1, v_2, \dots, v_t\} \in \mathbb{R}^n$  (all vectors are column vectors) where  $U, V$  are orthonormal sets and

$$u_i^T M = \sigma_i v_i \quad \text{and} \quad M v_i = \sigma_i u_i^T. \quad (1)$$

(When  $M$  is symmetric, each  $u_i = v_i$  and the  $\sigma_i$ 's are eigenvalues and can be negative.) Furthermore,  $M$  can be represented as

$$M = \sum_i \sigma_i u_i v_i^T. \quad (2)$$

The best rank  $k$  approximation to  $M$  consists of taking the first  $k$  terms of (2) and discarding the rest (where  $\sigma_1 \geq \sigma_2 \geq \dots \geq \sigma_r$ ).

Taking the best rank  $k$  approximation is also called *Principal Component Analysis* or PCA.

You probably have seen eigenvalue and eigenvector computations in your linear algebra course, so you know how to compute the PCA for symmetric matrices. The nonsymmetric case reduces to the symmetric one by using the following observation. If  $M$  is the matrix in (2) then

$$MM^T = \left( \sum_i \sigma_i u_i v_i^T \right) \left( \sum_i \sigma_i v_i u_i^T \right) = \sum_i \sigma_i^2 u_i u_i^T \quad \text{since } v_i^T v_j = 1 \text{ iff } i = j \text{ and } 0 \text{ else.}$$

Thus we can recover the  $u_i$ 's and  $\sigma_i$ 's by computing the eigenvalues and eigenvectors of  $MM^T$ , and then recover  $v_i$  by using (1).

Another application of singular vectors is the *Pagerank* algorithm for ranking webpages.

## 1.1 The power method

The eigenvalue computation you saw in your linear algebra course takes at least  $n^3$  time. Often we are only interested in the top few eigenvectors, in which case there's a method that can work much faster (especially when the matrix is *sparse*, i.e., has few nonzero entries).

As usual, we first look at the subcase of symmetric matrices. To compute the largest eigenvector of matrix  $M$  we do the following. Pick a random unit vector  $x$ . Then repeat the following a few times: replace  $x$  by  $Mx$ . We show this works under the following *Gap assumption*: There is a *gap* of  $\gamma$  between the the top two eigenvalues:  $|\lambda_1| - |\lambda_2| = \gamma$ .

The analysis is the same calculation as the one we used to analyse Markov chains. We can write  $x$  as  $\sum_i \alpha_i e_i$  where  $e_i$ 's are the eigenvectors and  $\lambda_i$ 's are numbered in decreasing order by absolute value. Then  $t$  iterations produces  $M^t x = \sum_i \alpha_i \lambda_i^t e_i$ . Since  $x$  is a unit vector,  $\sum_i \alpha_i^2 = 1$ .

Since  $|\lambda_i| \leq |\lambda_1| - \gamma$  for  $i \geq 2$ , we have

$$\sum_{i \geq 2} |\alpha_i| |\lambda_i^t| \leq n \alpha_{\max} (|\lambda_1| - \gamma)^t = n |\lambda_1|^t (1 - \gamma/|\lambda_1|)^t,$$

where  $\alpha_{\max}$  is the largest coefficient in magnitude.

Furthermore, since  $x$  was a random unit vector (and recalling that its projection  $\alpha_1$  on the fixed vector  $e_1$  is normally distributed), the probability is at least 0.99 that  $\alpha_1 > 1/(10n)$ . Thus setting  $t = O(\log n |\lambda_1|/\gamma)$  the components for  $i \geq 2$  become miniscule and  $x \approx \alpha_1 |\lambda_1|^t e_1$ . Thus rescaling to make it a unit vector, we get  $e_1$  up to some error. Then we can project all vectors to the subspace perpendicular to  $e_1$  and continue with the process to find the remaining eigenvectors and eigenvalues.

This process works under the above gap assumption. What if the gap assumption does not hold? Say, the first 3 eigenvalues are all close together, and separated by a gap from the fourth. Then the above process ends up with some random vector in the subspace spanned by the top three eigenvectors. For real-life matrices the gap assumption often holds.

## 2 Recovering planted bisections

Now we return to the planted bisection problem, also introduced last time.

The observation in Figure 1 suggests that the adjacency matrix is close to a rank 2 matrix shown there: the block within  $S_1, S_2$  have value  $p$  in each entry; the blocks between  $S_1, S_2$  have  $q$  in each entry. This is rank 2 since it has only two distinct column vectors.

Now we sketch why the best rank-2 approximation to the adjacency matrix will more or less recover the planted bisection. Specifically, the idea is to find the rank 2 approximation; with very high probability its columns can be cleanly clustered into 2 clusters. This gives a grouping of the vertices into 2 groups as well, which turns out to be the planted bisection.

Why this works has to do with the properties of rank  $k$  approximations. First we define two norms of a matrix.

**DEFINITION 1 (FROBENIUS AND SPECTRAL NORM)** *If  $M$  is an  $n \times n$  matrix then its Frobenius norm  $|M|_F$  is  $\sqrt{\sum_{ij} M_{ij}^2}$  and its spectral norm  $|M|_2$  is the maximum value of  $|Mx|_2$  over all unit vectors  $x \in \mathbb{R}^n$ . (By Courant-Fisher, the spectral norm is also the highest*

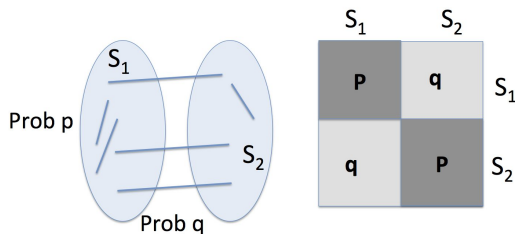


Figure 1: Planted Bisection problem: Edge probability is  $p$  within  $S_1, S_2$  and  $q$  between  $S_1, S_2$  where  $q < p$ . On the right hand side is the adjacency matrix. If we somehow knew  $S_1, S_2$  and grouped the corresponding rows and columns together, and squint at the matrix from afar, we'd see more density of edges within  $S_1, S_2$  and less density between  $S_1, S_2$ . Thus from a distance the adjacency matrix looks like a rank 2 matrix.

*eigenvalue.) For matrices that are not symmetric the definition of Frobenius norm is analogous and the spectral norm is the highest singular value.*

Last time we defined the *best rank  $k$  approximation to  $M$*  as the matrix  $\tilde{M}$  that is rank  $k$  and minimizes  $\left| M - \tilde{M} \right|_F^2$ . The following theorem shows that we could have defined it equivalently using spectral norm.

LEMMA 2

Matrix  $\tilde{M}$  as defined above also satisfies that  $\left| M - \tilde{M} \right|_2 \leq \left| M - B \right|_2$  for all  $B$  that have rank  $k$ .

THEOREM 3

If  $\tilde{M}$  is the best rank- $k$  approximation to  $M$ , then for every rank  $k$  matrix  $C$ :

$$\left| \tilde{M} - C \right|_F^2 \leq 5k \left| M - C \right|_2^2.$$

PROOF: Follows by Spectral decomposition and Courant-Fisher theorem, and the fact that the column vectors in  $\tilde{M}$  and  $C$  together span a space of dimension at most  $2k$ . Thus  $\left| \tilde{M} - C \right|_F^2$  involves a matrix of rank at most  $2k$ . Rest of the details are cut and pasted from Hopcroft-Kannan in Figure 2.

□

Returning to planted graph bisection, let  $M$  be the adjacency matrix of the graph with planted bisection. Let  $C$  be the rank-2 matrix that we *think* is a good approximation to  $M$ , namely, the one in Figure 1. Let  $\tilde{M}$  be the true rank 2 approximation found via SVD. In general  $\tilde{M}$  is not the same as  $C$ . But Theorem 3 implies that we can upper bound the average coordinate-wise squared difference of  $\tilde{M}$  and  $C$  by the quantity on the right hand side, which is the spectral norm (i.e., largest eigenvalue) of  $M - C$ .

Notice,  $M - C$  is a *random matrix* whose each coordinate is one of four values  $1 - p, -p, 1 - q, -q$ . More importantly, the expectation of each coordinate is 0 (since the entry of  $M$  is a coin toss whose expected value is the corresponding entry of  $C$ ). The study of eigenvalues of such random matrices is a famous subfield of science with unexpected connections to number theory (including the famous Riemann hypothesis), quantum physics (quantum gravity, quantum chaos), etc. We show below that  $\|M - C\|_2^2$  is at most  $O(np)$ . We conclude that the average column vector in  $\tilde{M}$  and  $C$  (whose square norm is about  $np$ ) are apart by  $O(p)$ . Thus intuitively, clustering the columns of  $C$  into two will find us the bipartition. Actually showing this requires more work which we will not do.

Here is a generic clustering algorithm into two clusters: Pick a random column of  $\tilde{M}$  and put into one cluster all columns whose distance from it is at most  $10p$ . Put all other columns in the other cluster.

## 2.1 Eigenvalues of random matrices

We sketch a proof of the following classic theorem to give a taste of this beautiful area.

THEOREM 4

Let  $R$  be a random matrix such that  $R_{ij}$ 's are independent random variables in  $[-1, 1]$  of expectation 0 and variance at most  $\sigma^2$ . Then with probability  $1 - \exp(-n)$  the largest eigenvalue of  $R$  is at most  $O(\sigma\sqrt{n})$ .

For simplicity we prove this for  $\sigma = 1$ .

PROOF: Recalling that the largest eigenvalue is  $\max_x |x^T R x|$ , we break the proof as follows.

Idea 1) For any fixed unit vector  $x \in \mathbb{R}^n$ ,  $|x^T R x| \leq O(\sqrt{n})$  with probability  $1 - \exp(-Cn)$  where  $C$  is an arbitrarily large constant. This follows from Chernoff-type bounds. Note that  $x^T R x = \sum_{ij} R_{ij} x_i x_j$ . By Chernoff bounds (Hoeffding's inequality) the probability that this exceeds  $t$  is at most

$$\exp\left(-\frac{t^2}{\sum_i x_i^2 x_j^2}\right) \leq \exp(-\Omega(t^2)),$$

since  $(\sum_{ij} x_i^2 x_j^2)^{1/2} \leq \sum_i x_i^2 = 1$ .

Idea 2) There is a set of  $\exp(n)$  special directions  $x_{(1)}, x_{(2)}, \dots$ , that approximately "cover" the set of unit vectors in  $\mathbb{R}^n$ . Namely, for every unit vector  $v$ , there is at least one  $x_{(i)}$  such that  $\langle v, x_{(i)} \rangle > 0.9$ .

First, note that  $\langle v, x_{(i)} \rangle > 0.9$  iff

$$|v - x_{(i)}|^2 = |v|^2 + |x_{(i)}|^2 - 2\langle v, x_{(i)} \rangle \leq 0.2.$$

In other words we are trying to cover the unit sphere with spheres of radius 0.2.

Try to pick this set greedily. Pick  $x_{(1)}$  arbitrarily, and throw out the unit sphere of radius 0.2 around it. Then pick  $x_{(2)}$  arbitrarily out of the remaining sphere, and throw out the unit sphere of radius 0.2 around it. And so on.

How many points did we end up with? By construction, each point that was picked has distance at least 0.2 from every other point that was picked, so the spheres of radius 0.1 around the picked points are mutually disjoint. Thus the maximum number of points

we could have picked is the number of disjoint spheres of radius 0.1 in a ball of radius at most 1.1. Denoting by  $B(r)$  denote the volume of spheres of volume  $r$ , this is at most  $B(1.1)/B(0.1) = \exp(n)$ .

Idea 3) Combining Ideas 1 and 2, and the union bound, we have with high probability,  $\left| x_{(i)}^T R x_{(i)} \right| \leq O(\sqrt{n})$  for all the special directions.

Idea 4): *If  $v$  is the eigenvector corresponding to the largest eigenvalue satisfies then there is some special direction satisfying  $\left| x_{(i)}^T R x_{(i)} \right| > 0.4v^T Rv$ .*

By the covering property, there is some special direction  $x_{(i)}$  that is close to  $v$ . Represent it as  $\alpha v + \beta u$  where  $u \perp v$  and  $u$  is a unit vector. So  $\alpha \geq 0.9$  and  $\beta \leq \sqrt{0.19} \leq 0.5$ . Then  $\left| x_{(i)}^T R x_{(i)} \right| = \alpha v^T R v + \beta u^T R u$ . But  $v$  is the largest eigenvalue so  $\left| u^T R u \right| \leq v^T R v$ . We conclude  $\left| x_{(i)}^T R x_{(i)} \right| \geq (0.9 - 0.5)v^T R v$ , as claimed.

The theorem now follows from Idea 3 and 4.  $\square$

#### BIBLIOGRAPHY

1. F. McSherry. *Spectral partitioning of random graphs*. In IEEE FOCS 2001 Proceedings.

**Lemma 8.7** Suppose  $A$  is an  $n \times d$  matrix and suppose  $C$  is an  $n \times d$  rank  $k$  matrix. Let  $\bar{A}$  be the best rank  $k$  approximation to  $A$  found by SVD. Then,  $\|\bar{A} - C\|_F^2 \leq 5k\|A - C\|_2^2$ .

**Proof:** Let  $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_k$  be the top  $k$  singular vectors of  $A$ . Extend the set of the top  $k$  singular vectors to an orthonormal basis  $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_p$  of the vector space spanned by the rows of  $\bar{A}$  and  $C$ . Note that  $p \leq 2k$  since  $\bar{A}$  is spanned by  $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_k$  and  $C$  is of rank at most  $k$ . Then,

$$\|\bar{A} - C\|_F^2 = \sum_{i=1}^k |(\bar{A} - C)\mathbf{u}_i|^2 + \sum_{i=k+1}^p |(\bar{A} - C)\mathbf{u}_i|^2.$$

Since  $\{\mathbf{u}_i | 1 \leq i \leq k\}$  are the top  $k$  singular vectors of  $A$  and since  $\bar{A}$  is the rank  $k$  approximation to  $A$ , for  $1 \leq i \leq k$ ,  $A\mathbf{u}_i = \bar{A}\mathbf{u}_i$  and thus  $|(\bar{A} - C)\mathbf{u}_i|^2 = |(A - C)\mathbf{u}_i|^2$ . For  $i > k$ ,  $\bar{A}\mathbf{u}_i = 0$ , thus  $|(\bar{A} - C)\mathbf{u}_i|^2 = |C\mathbf{u}_i|^2$ . From this it follows that

$$\begin{aligned} \|\bar{A} - C\|_F^2 &= \sum_{i=1}^k |(A - C)\mathbf{u}_i|^2 + \sum_{i=k+1}^p |C\mathbf{u}_i|^2 \\ &\leq k\|A - C\|_2^2 + \sum_{i=k+1}^p |A\mathbf{u}_i + (C - A)\mathbf{u}_i|^2 \end{aligned}$$

Using  $|a + b|^2 \leq 2|a|^2 + 2|b|^2$

$$\begin{aligned} \|\bar{A} - C\|_F^2 &\leq k\|A - C\|_2^2 + 2 \sum_{i=k+1}^p |A\mathbf{u}_i|^2 + 2 \sum_{i=k+1}^p |(C - A)\mathbf{u}_i|^2 \\ &\leq k\|A - C\|_2^2 + 2(p - k - 1) \sigma_{k+1}^2(A) + 2(p - k - 1) \|A - C\|_2^2 \end{aligned}$$

Using  $p \leq 2k$  implies  $k > p - k - 1$

$$\|\bar{A} - C\|_F^2 \leq k\|A - C\|_2^2 + 2k\sigma_{k+1}^2(A) + 2k\|A - C\|_2^2. \quad (8.1)$$

As we saw in Chapter 4, for any rank  $k$  matrix  $B$ ,  $\|A - B\|_2 \geq \sigma_{k+1}(A)$  and so  $\sigma_{k+1}(A) \leq \|A - C\|_2$  and plugging this in, we get the Lemma.  $\blacksquare$

Figure 2: Proof of Theorem 3 from Hopcroft-Kannan book